

# **A POLYNOMIAL-TIME APPROXIMATION ALGORITHM FOR ALL-TERMINAL NETWORK RELIABILITY**

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# The complexity of computing quantities

**P**: polynomial-time computable

**NP**: polynomial-time verifiable

Complexity class **#P** by Valiant (1979):

a counting analogue of **NP**.

Evaluation of probabilities;

Partition functions in statistical mechanics;

Counting discrete structures ...



## The complexity of approximate counting

What about (multiplicatively) approximating **#P**-complete problems?

- at most **NP**-hard (Stockmeyer, 1985);
- any polynomial approximation can be amplified into an  $(1 \pm \epsilon)$ -approximation with overhead polynomial in  $1/\epsilon$ .

Efficient approximation algorithms do exist! Most famous example: the permanent of a non-negative matrix (Jerrum, Sinclair, and Vigoda, 2004).

Many problems were proposed in 80s, and subsequently solved in 90s, but there are still a fair amount of leftovers!

# NETWORK RELIABILITY

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## Network reliability

Given a undirected graph (a.k.a. network)  $G = (V, E)$ , define a random subgraph  $G(p)$  by removing each edge independently with probability  $p$ .

One may be curious about all kinds of properties of  $G(p)$ :

- $\Pr[G(p) \text{ is connected}]$  (ALL-TERMINAL) RELIABILITY
- $\Pr[s \text{ and } t \text{ are connected in } G(p)]$   $s$ - $t$  RELIABILITY
- $\Pr[G(p) \text{ is acyclic}] \dots$

# Network reliability

(ALL-TERMINAL) RELIABILITY: The probability that  $G(p)$  is connected.

In other words, we want to compute

$$Z_{\text{rel}}(G, p) := \sum_{R \subseteq E: (V, R) \text{ is connected}} p^{|E \setminus R|} (1-p)^{|R|}.$$

For example:

$$Z_{\text{rel}}(\text{---} \circ \text{---} \circ \text{---} \circ \text{---} \circ \text{---}, p) = \text{---} \circ \text{---} \circ \text{---} \circ \text{---} \circ \text{---} = (1-p)^{n-1};$$

$$\begin{aligned} Z_{\text{rel}}(\square, p) &= \square + \square + \square + \square + \square \\ &= (1-p)^4 + 4p(1-p)^3; \end{aligned}$$

$$Z_{\text{rel}}(G, 1/2) = \frac{|\{R \subseteq E : (V, R) \text{ is connected}\}|}{2^{|E|}}.$$

## Computational complexity of reliability

Directed and undirected  $s$ - $t$  RELIABILITY (and a few other variants) are featured in the original list of 13 #P-complete problems by Valiant (1979).

Exact evaluation of ALL-TERMINAL RELIABILITY is shown to be #P-complete by Jerrum (1981), and independently Provan and Ball (1983).

What about approximation? Open since 80s.

Karger (1999) has given a famous FPRAS for UNRELIABILITY (namely  $1 - Z_{rel}$ ). However, approximating  $1 - Z_{rel}$  does not yield a good approximation for  $Z_{rel}$  when  $Z_{rel}$  is exponentially small.

# The Tutte polynomial

For a **connected** undirected graph  $G = (V, E)$ ,

$$Z_{\text{Tutte}}(G; x, y) := \sum_{R \subseteq E} (x-1)^{\kappa(R)-1} (y-1)^{\kappa(R)+|R|-|V|},$$

where  $\kappa(R)$  is the number of connected components of  $(V, R)$ .

A few specializations of  $(x, y)$ :

- $(1, 1)$ : # of spanning trees;
- $(1, 2)$ : # of connected subgraphs, and

$$Z_{\text{Tutte}}(G; 1, 1/p) = Z_{\text{rel}}(G, p) \cdot \frac{p^{|V|-|E|-1}}{(1-p)^{|V|-1}};$$

- $(2, 1)$ : # of forests (acyclic subgraphs);
- $(x-1)(y-1) = 2$  and  $x, y > 0$ : ferromagnetic Ising model.



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# Main result

Let  $m := |E|$  and  $n := |V|$ .

## Theorem (G. and Jerrum, 2018)

*There is a randomized algorithm approximating  $Z_{\text{rel}}$  within multiplicative factor  $(1 \pm \varepsilon)$ , with expected running time  $O(\varepsilon^{-2}(1-p)^{-3}m^2n^3)$ .*

## Theorem (G. and Jerrum, 2018)

*There is an exact sampler to draw (edge-weighted) connected subgraphs with expected running time  $O((1-p)^{-1}m^2n)$ .*

# **NATURAL ATTEMPTS**

(AND WHY THEY DO NOT SUCCEED)

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## Naive Monte Carlo

A natural unbiased estimator  $\tilde{Z}$  of  $Z_{\text{rel}}$ :

1. Draw  $k$  independent subgraphs  $(R_i)_{i \in [k]}$  of  $G(p)$ .
2. Let

$$\tilde{Z} := \frac{1}{k} \sum_{i \in [k]} \mathbb{1}_{\text{conn}}(R_i),$$

where  $\mathbb{1}_{\text{conn}}(R)$  is the indicator variable of  $(V, R)$  being connected.

It is easy to see that  $\mathbb{E} \tilde{Z} = Z_{\text{rel}}$ .

However, if  $Z_{\text{rel}}$  is exponentially small (e.g.  $Z_{\text{rel}}(P_n, p) = (1-p)^{n-1}$ ), then we will almost never see a connected  $R_i$ .

When that happens, the variance of  $\mathbb{1}_{\text{conn}}(R)$  is exponentially large, and  $k$  has to be exponentially large to yield a good approximation.

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# Unreliability

Nonetheless, naive Monte Carlo (NMC) is the basic building block of the FPRAS by [Karger \(1999\)](#) for **UNRELIABILITY** (namely  $1 - Z_{\text{rel}}$ ).

Karger's algorithm has been subsequently refined by [Harris and Srinivasan \(2014\)](#), [Karger \(2016, 2017\)](#).

[Karger \(2017\)](#) is a recursive algorithm using NMC running in  $O(n^{2.87})$ .

- Run NMC, if  $p^c > 1/2$ , where  $c$  is the size of the min-cut.
- Otherwise, draw subgraphs  $H_1, H_2 \sim G(q)$  where  $q = 2^{-1/c} > p$ , and

$$\frac{1}{2} (Z_{\text{unrel}}(H_1, p/q) + Z_{\text{unrel}}(H_2, p/q))$$

is an unbiased estimator of  $Z_{\text{unrel}}(G, p)$ .

- Recursively estimate  $Z_{\text{unrel}}(H_i, p/q)$  for  $i = 1, 2$ .

Similar ideas, once again, fail on graphs as simple as a path for  $Z_{\text{rel}}$ .



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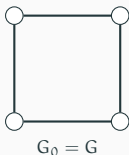
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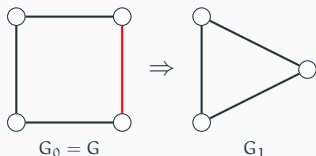
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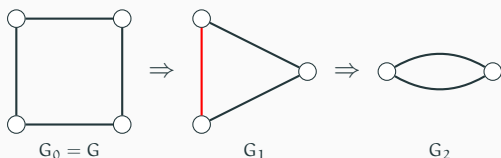
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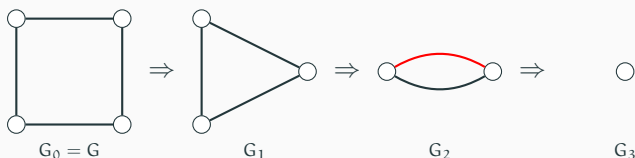




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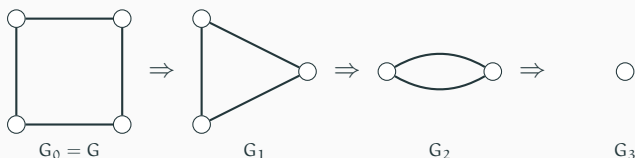
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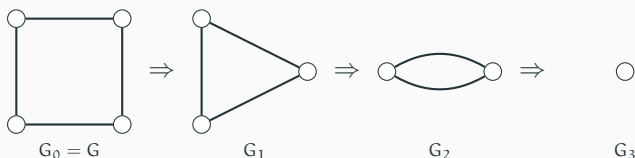
Rewrite

$$Z_{\text{rel}}(G) = \frac{Z_{\text{rel}}(G_0)}{Z_{\text{rel}}(G_1)} \cdot \frac{Z_{\text{rel}}(G_1)}{Z_{\text{rel}}(G_2)} \cdot \frac{Z_{\text{rel}}(G_2)}{Z_{\text{rel}}(G_3)} \cdot Z_{\text{rel}}(G_3).$$

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To estimate  $\frac{Z_{\text{rel}}(G_i)}{Z_{\text{rel}}(G_{i+1})}$ , draw  $C \sim \pi_{G_{i+1}}(\cdot)$  and let

$$C' := \begin{cases} C & \text{with prob. } p; \\ C \cup \{e\} & \text{otherwise,} \end{cases} \quad \text{and} \quad X := \mathbb{1}_{\text{conn}, G_i}(C').$$

Then  $\mathbb{E} X = \frac{Z_{\text{rel}}(G_i)}{Z_{\text{rel}}(G_{i+1})}$  and its variance is bounded by a polynomial.

# Markov chain Monte Carlo

There is a natural Markov chain converging to  $\pi_G(\cdot)$ :

1. Let  $C_0 = E$ .
2. Given  $C_t$ , randomly pick an edge  $e \in E$ .

If  $C_t \setminus \{e\}$  is disconnected then  $C_{t+1} = C_t$ . Otherwise,

$$C_{t+1} = \begin{cases} C_t \cup \{e\} & \text{with prob. } 1 - p; \\ C_t \setminus \{e\} & \text{with prob. } p. \end{cases}$$

Unfortunately, nothing is known about its mixing time (rate of convergence).

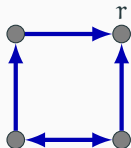
# **A SURPRISING EQUIVALENCE**

(AND AN ALTERNATIVE WAY TO SAMPLING)

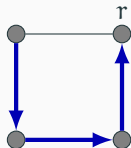
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# Reachability

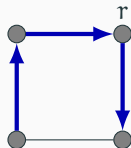
We say a directed graph  $D$  with root  $r$  is *root-connected* if all vertices can reach  $r$ .



Root-connected



Root-connected



Not root-connected!

**REACHABILITY:** in a directed graph  $D = (V, A)$  with root  $r$ , what's the probability that  $D(p)$  is root-connected?

$$Z_{\text{reach}}(D, p) := \sum_{R \subseteq A: (V, R) \text{ is root-connected}} p^{|\mathcal{A} \setminus R|} (1 - p)^{|R|}.$$

## A surprising equivalence

Ball (1980) showed that for any undirected graph  $G = (V, E)$ ,

$$Z_{\text{rel}}(G, p) = Z_{\text{reach}}(\vec{G}, p),$$

where  $\vec{G}$  is the directed graph obtained by replacing every  $e \in E$  with a pair of anti-parallel arcs. (Called **bi-directed**).



Thus we just need to approximate **REACHABILITY** in bi-directed graphs.

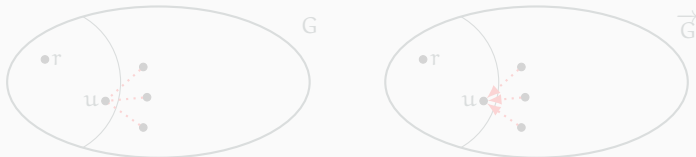
## A coupling proof

We have an alternative coupling proof of Ball's equivalence:

There is a coupling  $\mathcal{C}$  under which

$G(p)$  is connected  $\Leftrightarrow \vec{G}(p)$  is root-connected.

Explore  $G$  and  $\vec{G}$  like a BFS, starting from  $r$ . Reveal  $\vec{G}(p)$  and  $G(p)$  as the process proceeds. Couple the arc going towards the current vertex in  $\vec{G}(p)$  with the corresponding edge in  $G(p)$ .



When both exploration processes end, the sets of vertices that can reach  $r$  are exactly the same.



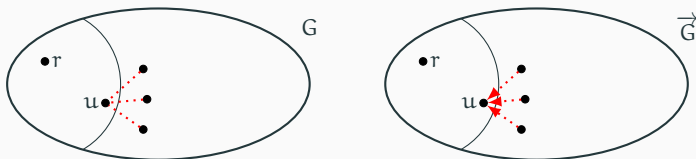
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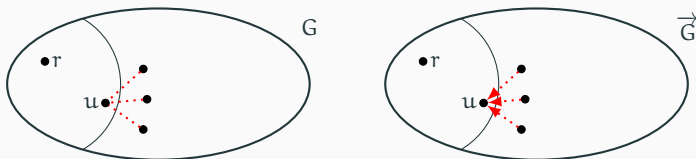
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# Cluster-popping

Goal: sample uniform (or edge-weighted) root-connected subgraphs.

Gorodezky and Pak (2014) proposed the “cluster-popping” algorithm:  
(Cluster: a subset of vertices not including  $r$  and with no arc going out.)

1. Let  $R$  be a subset of arcs by choosing each arc  $e$  with probability  $1 - p$  independently.
2. **While** there is at least one cluster in  $(V, R)$ :
  - Let  $C_1, \dots, C_k$  be all **minimal** clusters in  $(V, R)$ , and  $C = \bigcup_{i=1}^k C_i$ .
  - Re-randomize all arcs whose heads are in  $C$  to get a new  $R$ .

Gorodezky and Pak (2014) showed that this algorithm draws from the correct distribution, and they also conjectured that cluster-popping runs in expected polynomial time in bi-directed graphs.

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# Cluster-popping

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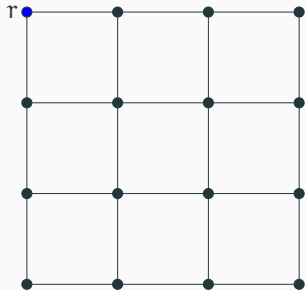
Gorodezky and Pak (2014) proposed the “cluster-popping” algorithm:  
(Cluster: a subset of vertices not including  $r$  and with no arc going out.)

1. Let  $R$  be a subset of arcs by choosing each arc  $e$  with probability  $1 - p$  independently.
2. **While** there is at least one cluster in  $(V, R)$ :
  - Let  $C_1, \dots, C_k$  be all **minimal** clusters in  $(V, R)$ , and  $C = \bigcup_{i=1}^k C_i$ .
  - Re-randomize all arcs whose heads are in  $C$  to get a new  $R$ .

Gorodezky and Pak (2014) showed that this algorithm draws from the correct distribution, and they also conjectured that cluster-popping runs in expected polynomial time in bi-directed graphs.

## An example run

Cluster-popping: repeatedly resample **minimal** clusters.

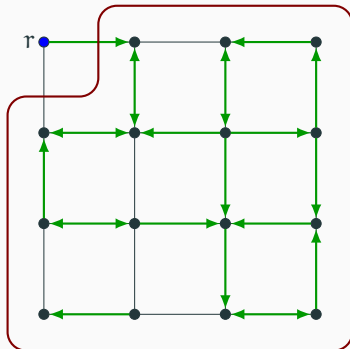






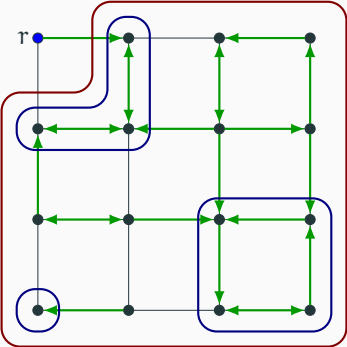
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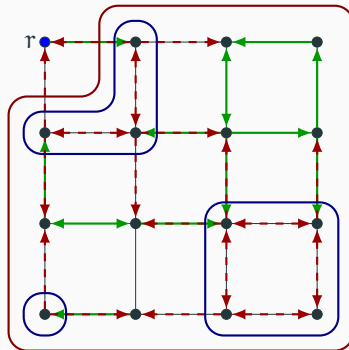
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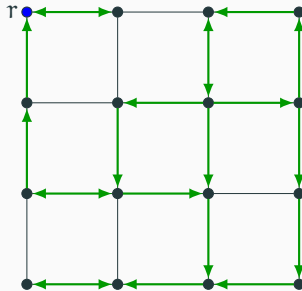
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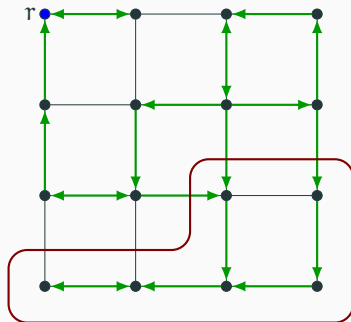
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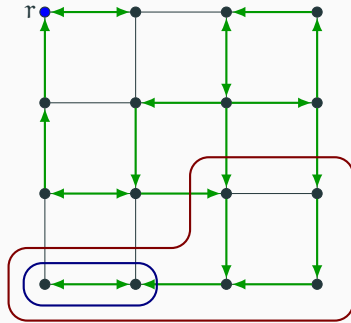
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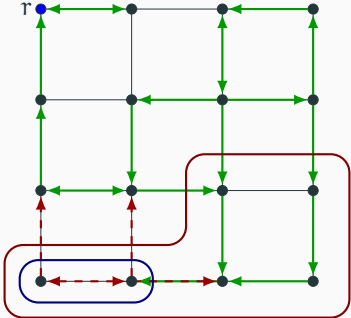
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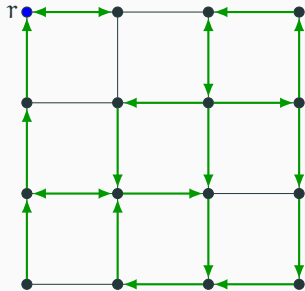
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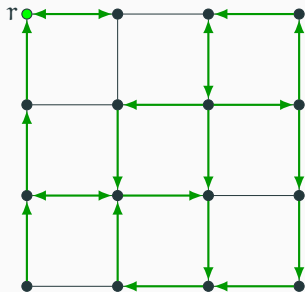
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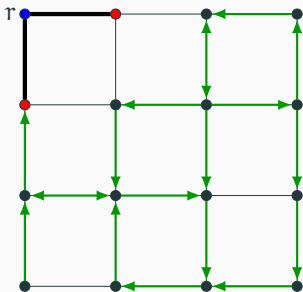
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Mapping back to connected subgraph.  
(Exploration order: left to right, bottom to top)

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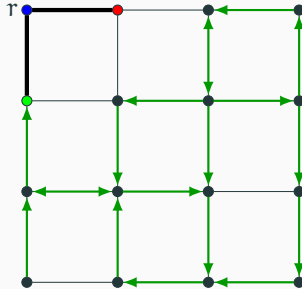
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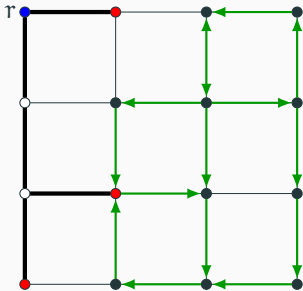


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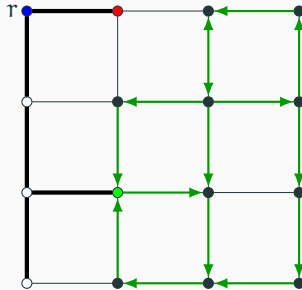


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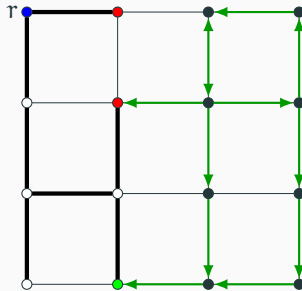
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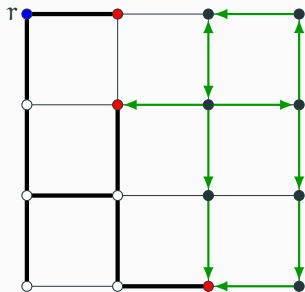
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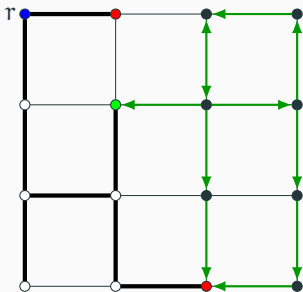
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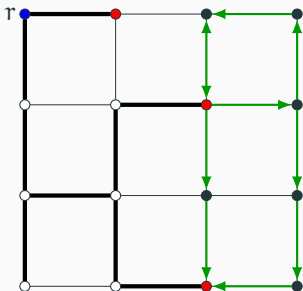
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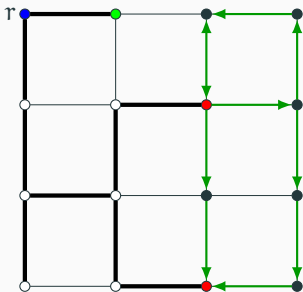
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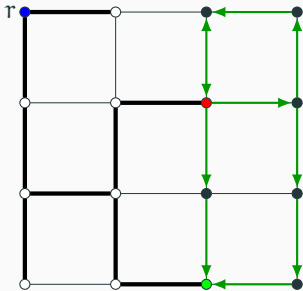
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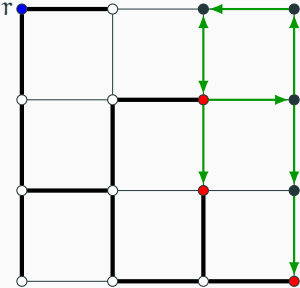
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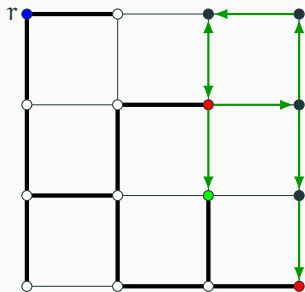


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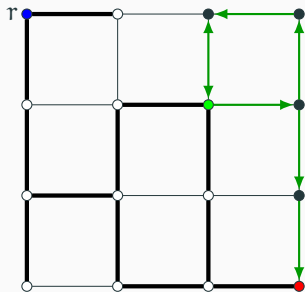
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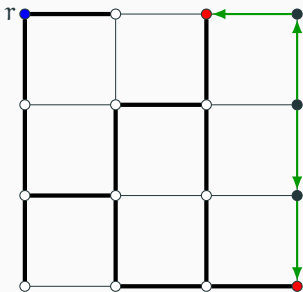
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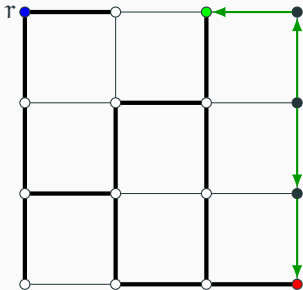
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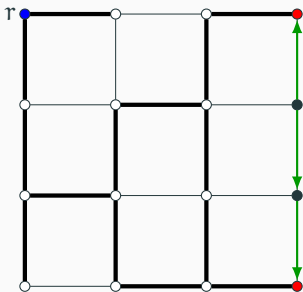
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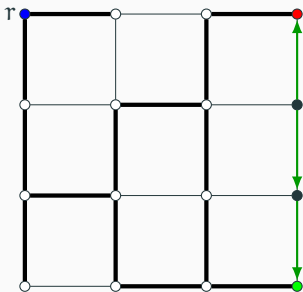
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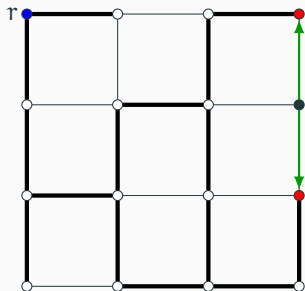
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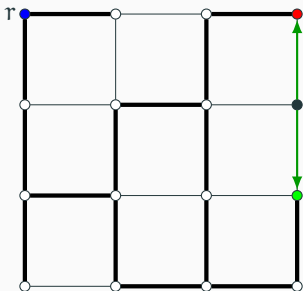
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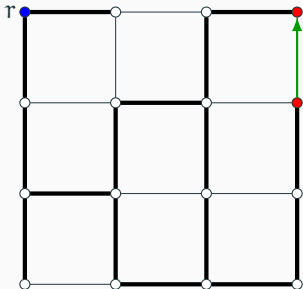


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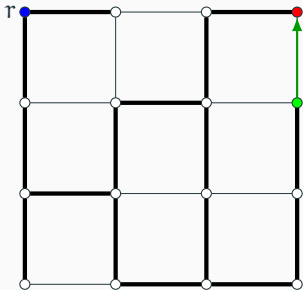
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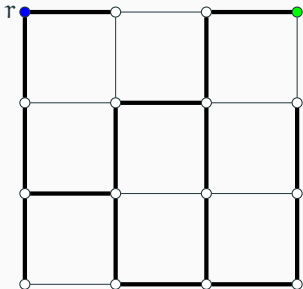
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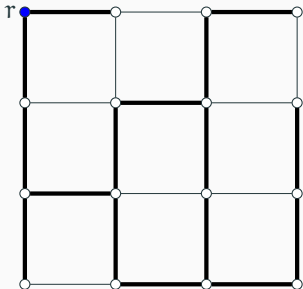
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# **PARTIAL REJECTION SAMPLING**

(A GENERAL THEORY BEHIND CLUSTER-POPPING)

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## Partial rejection sampling

Cluster-popping falls into the **PARTIAL REJECTION SAMPLING** framework (G., Jerrum, and Liu, 2017).

The goal is to sample from a product distribution, conditioned on a number of “bad” events not happening.

Rejection sampling throws away all variables.

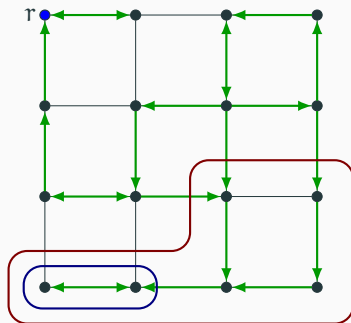
Instead, we want to recycle some randomness while resampling the “bad” events (and hopefully not too much more).

# Partial rejection sampling

Cluster-popping under partial rejection sampling:

Arcs are variables.

Minimal clusters are “bad” events.

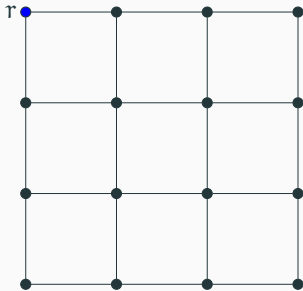


There can be exponentially many bad events.

## Extremal instances

An instance is called **extremal** (in the sense of [Shearer \(1985\)](#) regarding non-uniform Lovász Local Lemma):

if any two “bad” events  $A_i$  and  $A_j$  are either **independent** or **disjoint**.



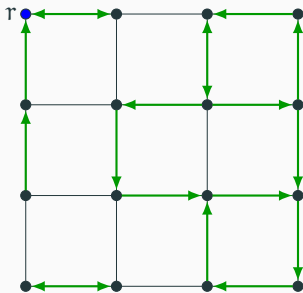
If the instance is **extremal**, then eliminating precisely the “**bad**” events in each iteration yields the correct distribution once halt ([GJL'17](#))!



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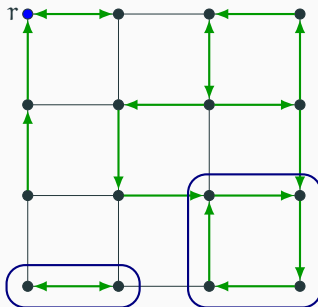


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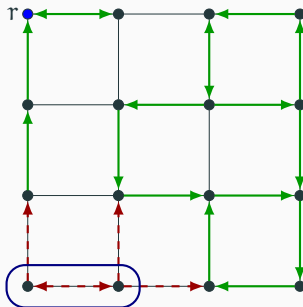


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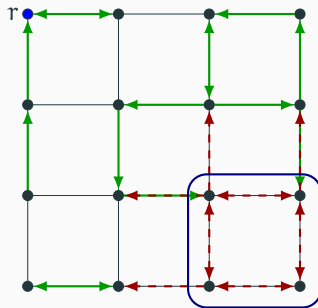


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## Resampling table

Associate an infinite stack  $X_{i,0}, X_{i,1}, \dots$  to each random variable  $X_i$ .  
When we need to resample, draw the next value in the stack.

|       |           |           |           |           |           |         |
|-------|-----------|-----------|-----------|-----------|-----------|---------|
| $X_1$ | $X_{1,0}$ | $X_{1,1}$ | $X_{1,2}$ | $X_{1,3}$ | $X_{1,4}$ | $\dots$ |
| $X_2$ | $X_{2,0}$ | $X_{2,1}$ | $X_{2,2}$ | $X_{2,3}$ | $X_{2,4}$ | $\dots$ |
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## Change the future, not the history

For **extremal** instances, replacing a **perfect** assignment with another one will not change the resampling history!

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| $X_2$ | $\Lambda_1$ |             | $X_{2,2}$ | $X_{2,3}$ | $X_{2,4}$ | $\dots$ |
| $X_3$ |             | $\Lambda_2$ |           | $X_{3,3}$ | $X_{3,4}$ | $\dots$ |
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| $X_2$ | $\Lambda_1$ | $X'_{2,1}$  | $X_{2,2}$  | $X_{2,3}$ | $X_{2,4}$ | $\dots$ |
| $X_3$ |             | $\Lambda_2$ | $X'_{3,2}$ | $X_{3,3}$ | $X_{3,4}$ | $\dots$ |
| $X_4$ |             | $X'_{4,1}$  | $X_{4,2}$  | $X_{4,3}$ | $X_{4,4}$ | $\dots$ |

## Change the future, not the history

For **extremal** instances, replacing a **perfect** assignment with another one will not change the resampling history!

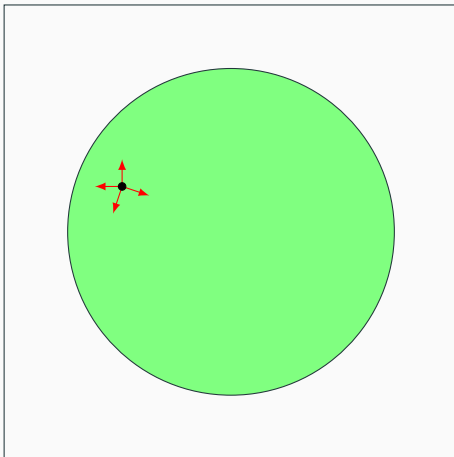
|       |             |             |            |           |           |         |
|-------|-------------|-------------|------------|-----------|-----------|---------|
| $X_1$ | $X'_{1,0}$  | $X_{1,1}$   | $X_{1,2}$  | $X_{1,3}$ | $X_{1,4}$ | $\dots$ |
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For any output  $\sigma$  and  $\tau$ , there is a **bijection** between trajectories leading to  $\sigma$  and  $\tau$ .

## Partial Rejection Sampling vs Markov chains

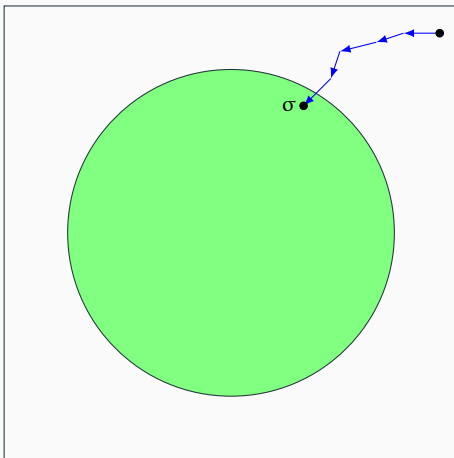
Markov chain is a random walk in the solution space.

(The solution space has to be connected,  
and the mixing time is not easy to analyze.)



## Partial Rejection Sampling vs Markov chains

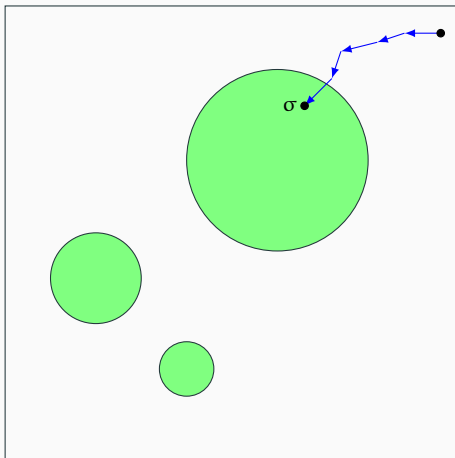
PRS is a local search on the **whole** space.





## Partial Rejection Sampling vs Markov chains

PRS is a local search on the **whole** space.  
(Ergodicity is not an issue.)

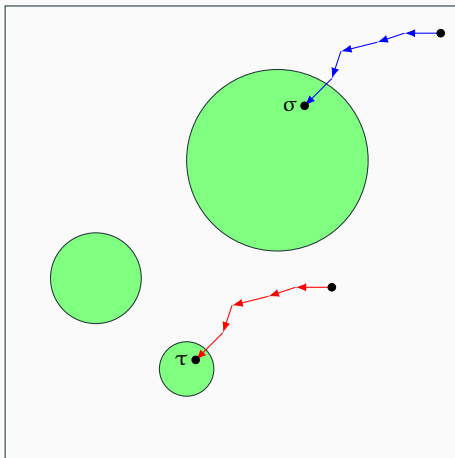


## Partial Rejection Sampling vs Markov chains

PRS is a local search on the **whole** space.

(Correctness guaranteed by the bijection.)

Exact formula for its running time on extremal instances.)



### Theorem (G., Jerrum, and Liu, 2017)

Under Shearer's condition, for *extremal* instances,

$$\mathbb{E} T = \frac{\text{total weight of one-flaw assignments}}{\text{total weight of perfect assignments}}.$$

(Shearer (1985) has shown a sufficient condition to guarantee the existence of one perfect assignment, which is optimal for Lovász Local Lemma.)

The upper bound is shown by Kolipaka and Szegedy (2011).

## Back to cluster-popping

Cluster-popping: repeatedly resample minimal clusters.

Let  $\Omega_k$  be the set of subgraphs with  $k$  minimal clusters, and

$$Z_k := \sum_{S \in \Omega_k} p^{|\mathbb{E} \setminus S|} (1-p)^{|S|}. \quad \text{Then, } \mathbb{E} T = \frac{Z_1}{Z_0}.$$

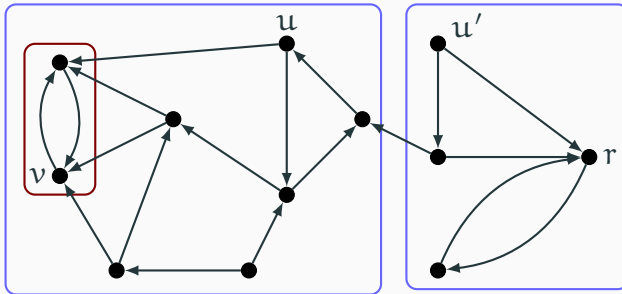
**Lemma** (G. and Jerrum, 2018)

For *bi-directed* graphs,  $Z_1 \leq \frac{p}{1-p} \cdot \text{mn} Z_0$ .

We show this by designing an injective mapping  $\Omega_1 \rightarrow \Omega_0 \times V \times E$ .

# Injective mapping

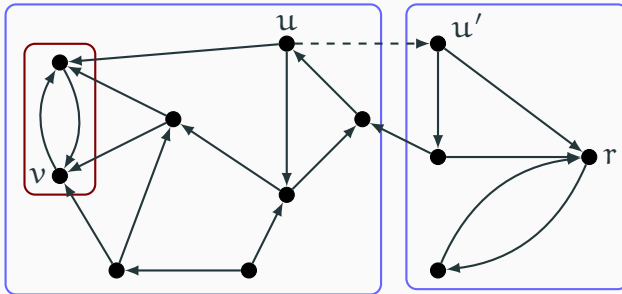
Given  $R \in \Omega_1$ , we map it to  $R_0 \in \Omega_0$  by “repairing” the unique minimal cluster.



Conversely, given  $R_0 \in \Omega_0$ ,  $(u, u')$  and  $v$ , we can recover  $R \in \Omega_1$ .

# Injective mapping

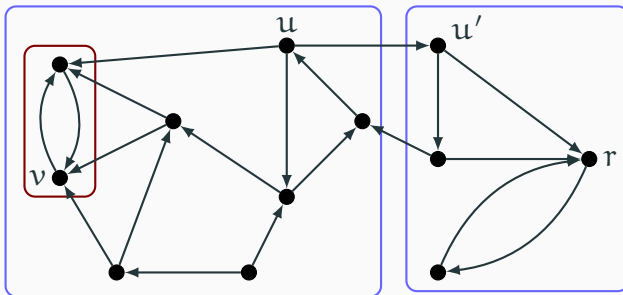
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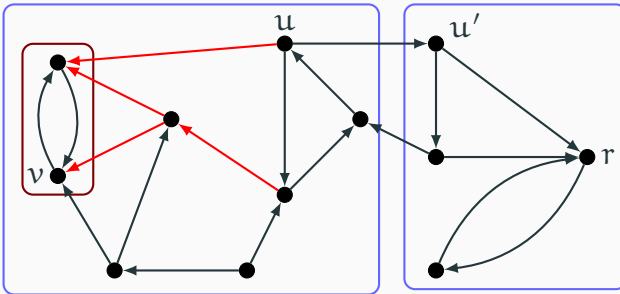
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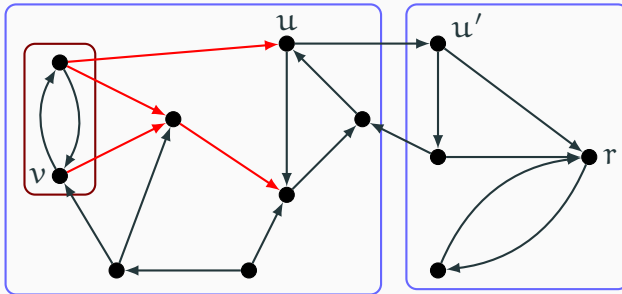


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# Injective mapping

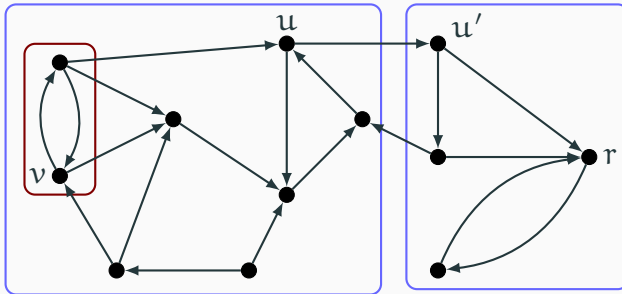
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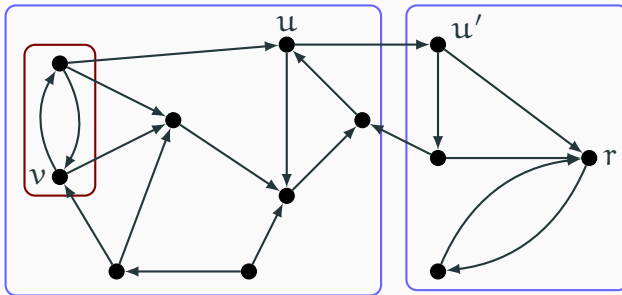
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## Recap for reliability

Approximate  $Z_{\text{rel}}(G)$  via a sequence of contractions  $G_0, \dots, G_{n-1}$ , and estimate each  $\frac{Z_{\text{rel}}(G_i)}{Z_{\text{rel}}(G_{i+1})}$  using the following sampling oracle:

1. run cluster-popping to sample a root-connected subgraph in  $\vec{G}$ ;
2. use the coupling to get a random connected subgraph.

To bound the running time of cluster-popping, we use a result of (GJL'17) and design an injective mapping.

## Counting connected subgraphs of fixed size

Let  $S_t$  be the set of connected subgraph of size  $t$  where  $n - 1 \leq t \leq m$  and  $N_t = |S_t|$ . Then a result of [Huh and Katz \(2012\)](#) implies that the sequence  $(N_t)_t$  is log-concave, namely,

$$N_{t-1}N_{t+1} \leq N_t^2.$$

(Here we treat connected subgraphs as independent sets in the dual of the graphic matroid, which is representable and [HK'12](#) applies. Similar log-concavity in general matroid is resolved by [Adiprasito, Huh, and Katz \(2015\)](#).)

Given the sampler for connected subgraphs and log-concavity, we can set  $\mathbf{p} = \frac{N_t}{N_{t-1} + N_t}$  so that subgraphs in  $S_t$  show up frequently enough. There is a standard approach ([Jerrum and Sinclair, 1989](#)) to estimate each individual  $N_t$ .

## Other examples of PRS

Extremal instances:

- Uniform spanning trees – cycle-popping (Wilson, 1996)
- Uniform sink-free orientations – sink-popping (Bubley and Dyer, 1997) (Cohn, Pemantle, and Propp, 2002)

General instances (G., Jerrum, and Liu, 2017):

- Weighted independent set (Hardcore gas model)
- Hard disks / hard spheres model (G. and Jerrum, 2018)
- Solutions to  $k$ -CNF formulas with bounded variable degrees

Results for general instances are far from optimal.

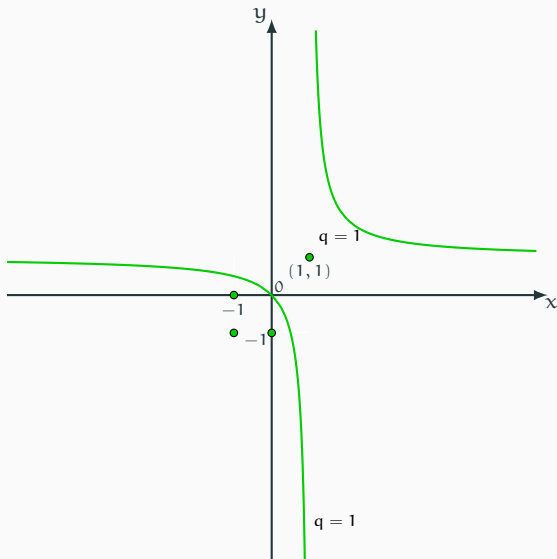
## **CONCLUDING REMARKS**

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# Approximating the Tutte polynomial

$$q = (x - 1)(y - 1)$$

Poly-time



Ref:

Jaeger, Vertigan, and Welsh (1990);

Jerrum and Sinclair (1993);

Goldberg and Jerrum (2008, 2012, 2014)

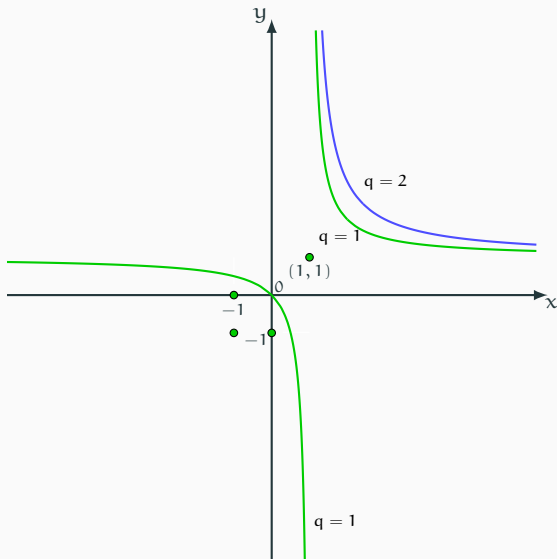


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(#P-hard mostly)

#PM-equivalent

#BIS-hard

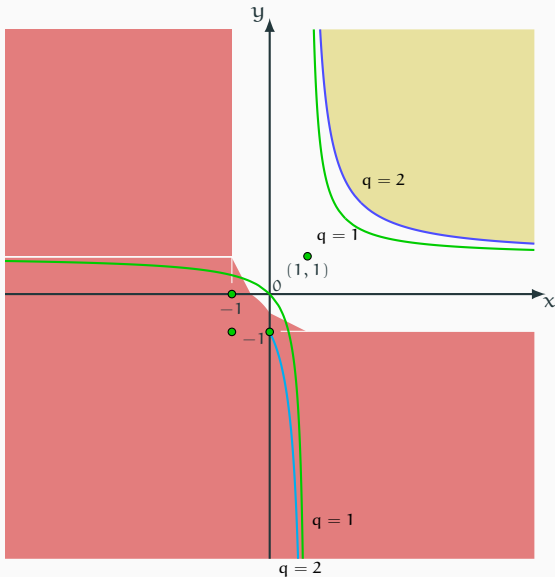
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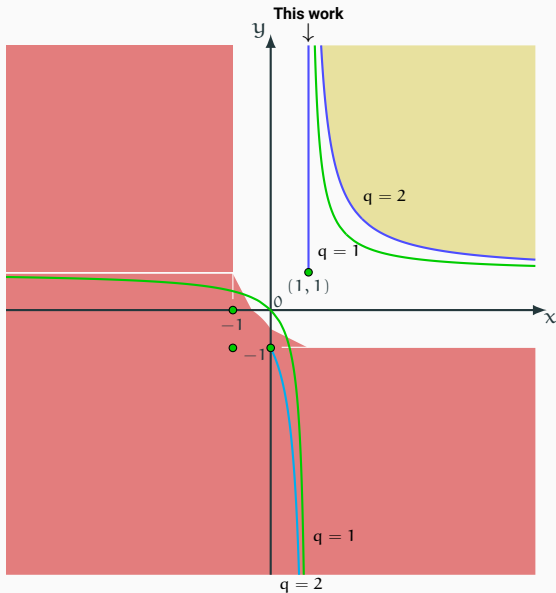
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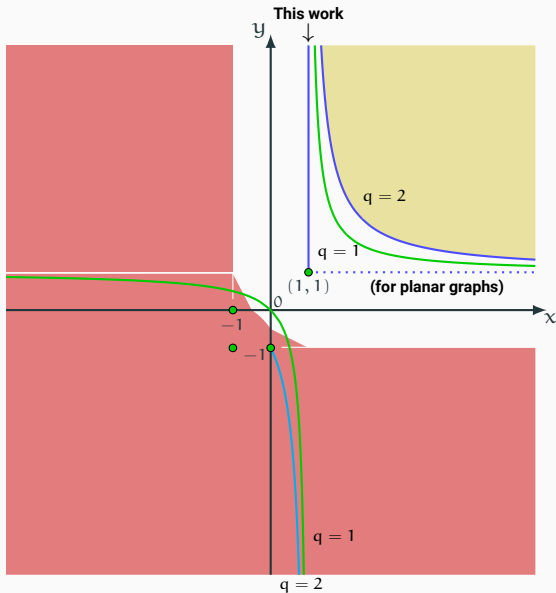
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## A common paradigm

Both our result and the previous positive result on the Tutte plane (Jerum and Sinclair, 1993) follow the same pattern:

1. Transform the problem into an equivalent one:
  - Ferromagnetic Ising model  $\rightarrow$  even subgraphs (JS'93);
  - Reliability  $\rightarrow$  bi-directed reachability.
2. Exploit some nice properties of the new solution space.

Are there other equivalences we have not discovered yet?

## Open problems

- Is the **Markov chain** for connected subgraphs rapidly mixing?
- Approximating **s-t RELIABILITY**, and other variants?  
(The natural Markov chain is exponentially slow for s-t version.)
- Approximating  $Z_{\text{Tutte}}(G; \chi, 1)$  for  $\chi > 1$  (edge-weighted forests)?

# THANK YOU!

arXiv:1611.01647  
(PARTIAL REJECTION SAMPLING)

arXiv:1709.08561  
(NETWORK RELIABILITY)