

AN FPRAS FOR TWO TERMINAL RELIABILITY IN DIRECTED ACYCLIC GRAPHS

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Based on joint work with **Weiming Feng** (ETH Zürich)

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THE COMPLEXITY OF COMPUTING QUANTITIES

Complexity class #P by Valiant (1979):

a counting analogue of NP

E.g. counting the number of solutions to CNF formulas

Other examples:

determinants / permanents of matrices;

evaluation of probabilities;

partition functions in statistical physics;

counting discrete structures ...



THE COMPLEXITY OF APPROXIMATE COUNTING

What about (multiplicatively) approximating #P-complete problems?

- at most NP-hard (Stockmeyer 1983; Valiant and Vazirani, 1986);
- for many problems, polynomial approximation can be amplified into ϵ -approximation paying only a polynomial runtime overhead.

Typically, an approximate counting problem is either NP-hard or FPRASable.

FPRASes do exist. Famous examples include:

- the number of solutions to DNF formulae (Karp and Luby, 1983);
- all-terminal network (un)reliability (Karger 1999);
- the volume of convex bodies (Dyer, Frieze, and Kannan, 1991);
- the permanent of non-negative matrices (Jerrum, Sinclair, and Vigoda, 2004).

However, there are still many open problems in approximate counting!

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NETWORK RELIABILITY



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Given a directed or undirected graph (a.k.a. network) $G = (V, E)$, define a random subgraph $G(p)$ by **removing** each edge independently with probability p .

TWO-TERMINAL RELIABILITY: given $s, t \in V$,

$$\Pr \left[s \xrightarrow{G(p)} t \right]$$

Directed and undirected **TWO-TERMINAL RELIABILITY** (and a few other variants) are featured in the original list of 13 #P-complete problems by Valiant (1979).

One may ask the probability of other properties of $G(p)$, such as whether $G(p)$ is connected (**ALL-TERMINAL RELIABILITY**).

NETWORK RELIABILITY

TWO-TERMINAL RELIABILITY: $\Pr \left[s \xrightarrow{G(p)} t \right]$

In other words, we want to compute

$$Z_{\text{rel}}(G, p) := \sum_{R \subseteq E: s \xrightarrow{(V, R)} t} p^{|\mathcal{E} \setminus R|} (1-p)^{|R|}$$

For example:

$$Z_{\text{rel}}(s \text{---} \dots \text{---} t, p) = s \text{---} \dots \text{---} t = (1-p)^{n-1}$$

$$Z_{\text{rel}}(s \square t, p) = \square + \square + \square + \square + \square + \square + \square$$

$$= (1-p)^4 + 4p(1-p)^3 + 2p^2(1-p)^2$$

COMPUTATIONAL COMPLEXITY OF RELIABILITY

Exact evaluation of almost all variants of reliability is $\#P$ -complete, shown by the pioneer work of Valiant (1979), Jerrum (1981), Provan and Ball (1983), etc.

Much less is known about their approximation complexity.

A famous breakthrough is by Karger (1999), who gave an FPRAS for **ALL-TERMINAL UNRELIABILITY**.

(However, good approximation of *unreliability* is not necessarily a good approximation for *reliability* when reliability is exponentially small.)

ALL-TERMINAL RELIABILITY is shown to have an FPRAS by G. and Jerrum (2019), resolving positively conjectures by Provan and Ball (1983), Welsh (1993), Karger (1999), etc.

MAIN RESULTS

We gave an FPRAS for **TWO-TERMINAL RELIABILITY** in directed acyclic graphs (DAGs).

Theorem (Feng and G., 2023)

Let G be a DAG and \mathbf{p} denote failure probabilities. There is a randomized algorithm that takes $(G, \mathbf{p}, s, t, \varepsilon)$ as inputs and outputs a $(1 \pm \varepsilon)$ -approximation to $s - t$ RELIABILITY with probability at least $3/4$ in time $\tilde{O}(n^6 m^4 \max\{m^4, \varepsilon^{-4}\})$.

This answers positively a conjecture by Zenklusen and Laumanns (2011).

On the flip side, the corresponding unreliability problem is unlikely to have an FPRAS.

Theorem (Feng and G., 2023)

There is no FPRAS to estimate $s - t$ UNRELIABILITY in DAGs unless there is an FPRAS for **#BIS**.

Here **#BIS** is the problem of counting independent sets in bipartite graphs. It is conjectured to have no FPRAS.

Independently, [Amarilli, van Bremen, and Meel \(2023\)](#) reduce $s-t$ *RELIABILITY* in DAGs to counting the number of accepting paths of a given length for non-deterministic automata (*#NFA*). The latter problem has an FPRAS by [Arenas, Croquevielle, Jayaram, and Riveros \(2021\)](#).

[ACJR21](#)'s algorithm runs in time $O\left(\left(\frac{n\ell}{\varepsilon}\right)^{17}\right)$ for an n -state NFA and strings of length ℓ .

[AvBM23](#) reduces a DAG with n vertices and m edges to a *#NFA* instance counting length m accepting strings for an NFA with $O(m^2)$ states. Thus their running time is like $O\left(\frac{m^{51}}{\varepsilon^{17}}\right)$.

COMPUTATIONAL COMPLEXITY OF APPROXIMATING RELIABILITY

Terminal	Graphs	Type	Complexity	Best run-time
All	Undirected	Unrel	FPRAS (K99)	$\frac{m^{1+o(1)}}{\varepsilon^2} + \tilde{O}\left(\frac{n^{1.5}}{\varepsilon^3}\right)$ (CHLP23)
All	Undirected	Rel	FPRAS (GJ19)	$\tilde{O}\left(\frac{mn}{\varepsilon^2}\right)$ (CGZZ23)
s – t	DAG	Rel	FPRAS (FG23, AvBM23)	$\tilde{O}(n^6 m^4 \max\{m^4, \varepsilon^{-4}\})$ (FG23)
s – t	DAG	Unrel	#BIS-hard (FG23)	—
S – t	DAG	Rel	NP-hard (upcoming)	—

CHLP23: [Cen, He, Li, and Panigrahi \(2023\)](#)

CGZZ23: [Chen, G., Zhang, and Zou \(2023\)](#)

SOME NATURAL ATTEMPTS

(AND WHY THEY DO NOT SUCCEED)



A natural unbiased estimator \tilde{Z} of Z_{rel} :

1. Draw k independent subgraphs $(S_i)_{i \in [k]}$ of $G(p)$.
2. Let

$$\tilde{Z} := \frac{1}{k} \sum_{i \in [k]} \mathbb{1}_{s \rightarrow t}(S_i),$$

where $\mathbb{1}_{s \rightarrow t}(S)$ is the indicator variable whether $s \xrightarrow{(V,S)} t$.

It is easy to see that $\mathbb{E} \tilde{Z} = Z_{\text{rel}}$.

However, if Z_{rel} is exponentially small (e.g. $Z_{\text{rel}}(P_n, p) = (1-p)^{n-1}$), then we will almost never see a connected S_i .

In that case, the relative variance of $\mathbb{1}_{s \rightarrow t}(S)$ is exponentially large, and k has to be exponentially large to yield a good approximation.

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Karger's method for (undirected) **ALL-TERMINAL UNRELIABILITY**:

Naive Monte Carlo + random contraction

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via the counting to sampling reduction (Jerrum, Valiant, and Vazirani, 1986) and then either

- Partial rejection sampling (G. and Jerrum, 2019); or
- Markov chain (Anari, Liu, Oveis Gharan, and Vintzant, 2019).

All these techniques rely on some nice structure of the solution space, which **TWO-TERMINAL RELIABILITY** (even in DAG) does not have.

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MARKOV CHAIN MONTE CARLO

Markov chains are the “off the shelf” approach to sampling from complicated distributions.

Here we want to sample $R \sim G(p)$ conditioned on $s \xrightarrow{R} t$.

There is a natural Markov chain converging to the desired distribution:

1. Let $C_0 = E$.
2. Given C_t , randomly pick an edge $e \in E$.

If $\neg \left(s \xrightarrow{C_t \setminus \{e\}} t \right)$ then $C_{t+1} = C_t$. Otherwise,

$$C_{t+1} = \begin{cases} C_t \cup \{e\} & \text{with prob. } 1 - p; \\ C_t \setminus \{e\} & \text{with prob. } p. \end{cases}$$

This chain is slow for s and t connected via two length $n/2$ paths.

**DYNAMIC PROGRAMMING
MEETS MONTE CARLO**

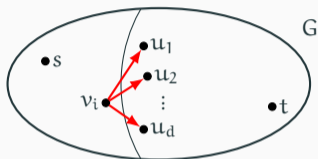


Our main inspiration is the #NFA algorithm of [Arenas, Croquevielle, Jayaram, and Riveros \(2021\)](#), which in turn builds upon a quasi-polynomial time algorithm for generating words in context-free grammars (#CFG) by [Gore, Jerrum, Kannan, Sweedyk, and Mahaney \(1997\)](#).

Given the topological ordering $s = v_1, v_2, \dots, v_n = t$ of the input DAG, we estimate $R_i := \Pr \left[v_i \xrightarrow{G(p)} t \right]$ from $i = n$ to 1.

Note that for v_i , we can safely ignore all vertices unreachable from v_i . Call the resulting graph G_{v_i} , and we only need to consider $G_{v_i}(p)$.

THE INDUCTIVE STEP



Let v_i 's neighbours be u_1, \dots, u_d . Then,

$$\Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \right] = \Pr \left[\exists j \in [d], ((v_i, u_j) \in G_{v_i}(p)) \wedge (u_j \xrightarrow{G_{u_j}(p)} t) \right].$$

This reminds us to use the DNF counting technique of [Karp and Luby \(1983\)](#).

Given DNF $\Phi = C_1 \vee C_2 \vee \dots \vee C_m$, counting the number of solutions to Φ is equivalent to evaluating

$$\Pr[\sigma \models \Phi] = \Pr[\exists i \in [m], \sigma \models C_i],$$

where σ is a uniformly at random assignment.

Let Ω_i be the set of solutions to C_i , and $p_i := \Pr[\sigma \models C_i] = \frac{|\Omega_i|}{2^n} = 2^{-|C_i|}$.

Karp-Luby goes as follows:

1. Draw i with probability proportional to p_i ;
2. draw uniformly σ such that $\sigma \models C_i$ (namely uniformly from Ω_i);
3. Let t_σ be the number of clauses σ satisfies. Output $Z := \frac{1}{t_\sigma}$.

Then $\mathbb{E}[Z] = \frac{1}{\sum_{i \in [m]} p_i} \cdot \Pr[\exists i \in [m], \sigma \models C_i]$.

Also, $\mathbb{E}[Z] \geq \frac{1}{m}$, implying that $\frac{\text{Var}(Z)}{\mathbb{E}[Z]^2} \leq m^2$.

KARP-LUBY IN ACTION (FOR RELIABILITY)

We want to estimate $\Pr \left[\exists j \in [d], ((v_i, u_j) \in G_{v_i}(p)) \wedge (u_j \xrightarrow{G_{u_j}(p)} t) \right] = \Pr [\exists j \in [d], \mathcal{E}_j]$.

Let $p_j := \Pr [\mathcal{E}_j] = \Pr [(v_i, u_j) \in G_{v_i}(p)] \times \Pr \left[u_j \xrightarrow{G_{u_j}(p)} t \right] = (1 - p)R_{u_j}$.

Then we can

1. draw $j \in [d]$ with probability proportional to p_j ;
2. draw $S \sim G_{u_j}(p)$ conditioned on $u_j \xrightarrow{S} t$; draw other edges in G_{v_i} independently with probability $1 - p$; then add (v_i, u_j) to get a sample S' ;
3. let $t_{S'}$ be the number of events \mathcal{E}_j occurring under S' ; output $Z := \frac{1}{t_{S'}}$.

The same analysis implies that the expectation of this estimator is what we want and the relative variance is small.

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ARE WE THERE YET?

To implement [Karp-Luby](#), we need to do two things:

- calculate $p_j = (1 - p)R_{u_j}$, which can be done as R_{u_j} is known in previous steps;
- sample $S \sim G_{u_j}(p)$ conditioned on $u_j \xrightarrow{S} t$.

How do we sample?

The sampling to counting reduction *a la* [Jerrum, Valiant, and Vazirani \(1986\)](#) to the rescue!

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SAMPLING TO COUNTING REDUCTION

JVV'86 self-reduction:

sample edges one-by-one, using marginal probabilities conditioned on previous outcomes



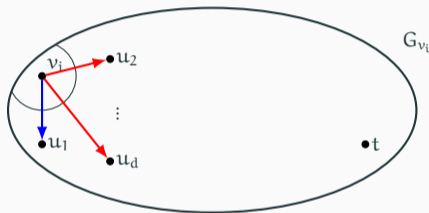
The marginal of $e = (v_i, u_1)$ is

$$\frac{\Pr[e \in G_{v_i}(p)] \Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \in G_{v_i}(p) \right]}{\Pr[e \in G_{v_i}(p)] \Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \in G_{v_i}(p) \right] + \Pr[e \notin G_{v_i}(p)] \Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \notin G_{v_i}(p) \right]}$$

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SAMPLING TO COUNTING REDUCTION

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The marginal of $e = (v_i, u_1)$ is

$$\frac{(1-p) \Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \in G_{v_i}(p) \right]}{(1-p) \Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \in G_{v_i}(p) \right] + p \Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \notin G_{v_i}(p) \right]}$$

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How do we estimate $\Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \in G_{v_i}(p) \right]$?



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Contract v_i and u_1 . Call it v_i' . Note that we would have already computed (R_u, S_u) for u that is a neighbour of v_i or u_1 (as it comes after v_i in the topological order).

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How do we estimate $\Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \in G_{v_i}(p) \right]$?



Contract v_i and u_1 . Call it v'_i . Note that we would have already computed (R_u, S_u) for u that is a neighbour of v_i or u_1 (as it comes after v_i in the topological order).

That's all [Karp-Luby](#) needs to estimate $\Pr \left[v'_i \xrightarrow{G_{v'_i}(p)} t \right]$!

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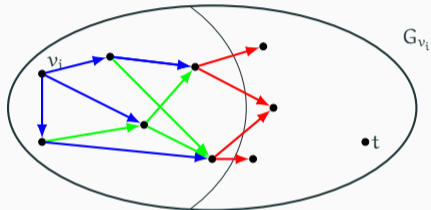
Contract v_i and u_1 . Call it v'_i . Note that we would have already computed (R_u, S_u) for u that is a neighbour of v_i or u_1 (as it comes after v_i in the topological order).

That's all [Karp-Luby](#) needs to estimate $\Pr \left[v'_i \xrightarrow{G_{v'_i}(p)} t \right]$!

The case for $\Pr \left[v_i \xrightarrow{G_{v_i}(p)} t \mid e \notin G_{v_i}(p) \right]$ is similar. Then move on to the next edge and repeat.

COMPUTING THE MARGINAL

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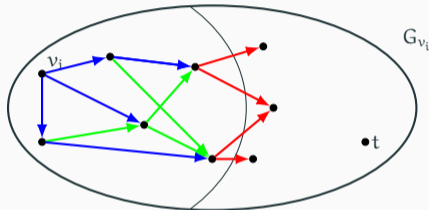
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- **Blue** edges: those chosen using the marginal by [Karp-Luby](#), forming a tree.
- **Green** edges: conditioned on the blue edges, they have no affect on reaching t , and thus have marginal $1 - p$.
- **Red** edges: the current frontier, whose tails have been processed (namely (R_u, S_u) has been computed already).

EFFICIENCY OF THE ALGORITHM

The self-reduction to generate one sample uses, say, k samples for each Karp-Luby step. If we generate fresh samples each time, in total at least k^n samples are required. The key to be efficient here is that samples can be reused!

Reusing samples introduces subtle correlation among R_{v_i} 's and the samples. However:

- accurate marginals $\xrightarrow{\text{union bound over all edges}}$ accurate samples;
- accurate samples $\xrightarrow{\text{union bound over all possibilities}}$ accurate marginals.

Accurate marginals come from accurate estimation of the R_{v_i} 's. The samples are subtly correlated, but they can be coupled with independent fresh samples with small error.

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HARDNESS RESULTS

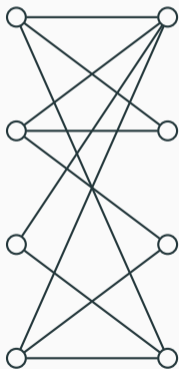


#BIS-HARDNESS FOR TWO-TERMINAL VERTEX UNRELIABILITY IN DAGs

#BIS

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TWO-TERMINAL VERTEX UNRELIABILITY

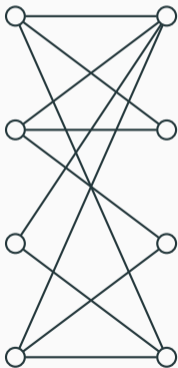


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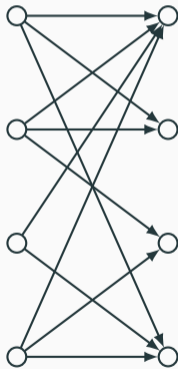
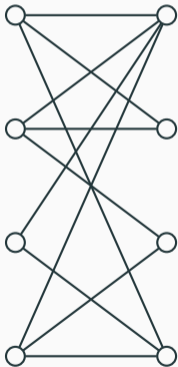


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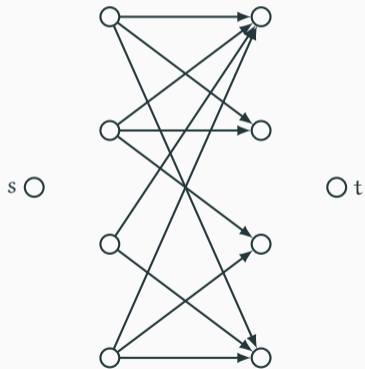
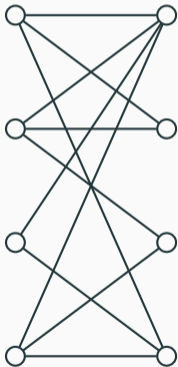


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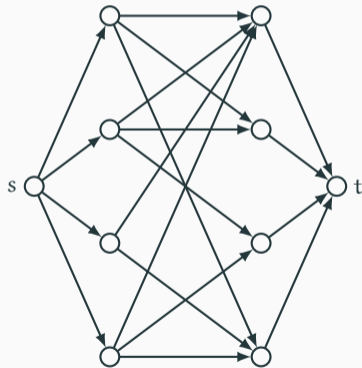
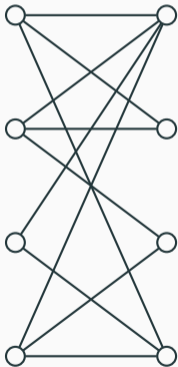


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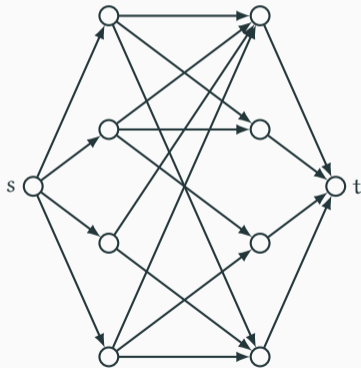
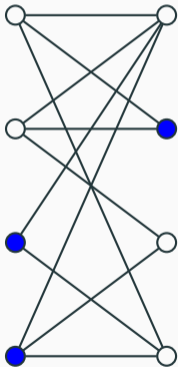


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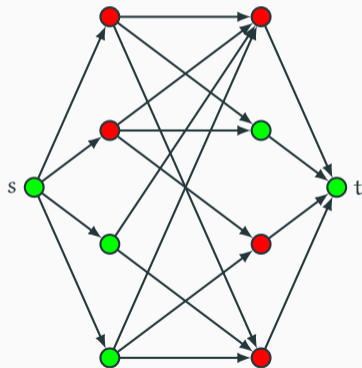
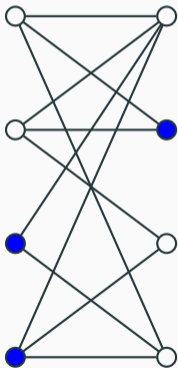


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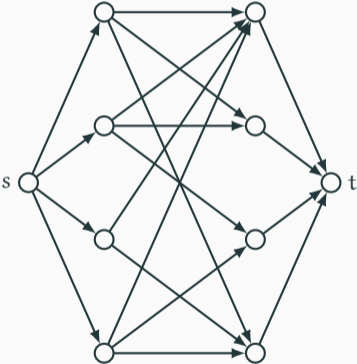
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TWO-TERMINAL VERTEX UNRELIABILITY



FROM VERTEX TO EDGE UNRELIABILITY IN DAGs

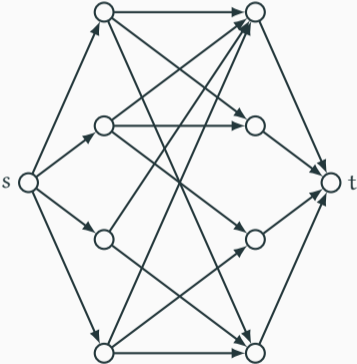
VERTEX UNRELIABILITY



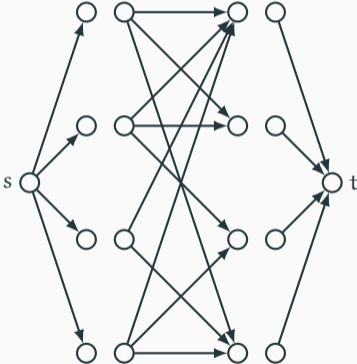
EDGE UNRELIABILITY

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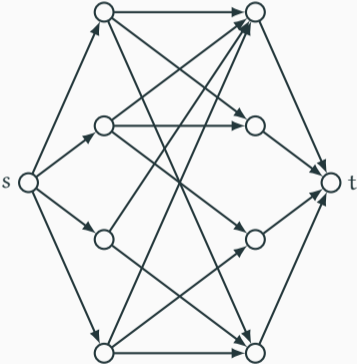


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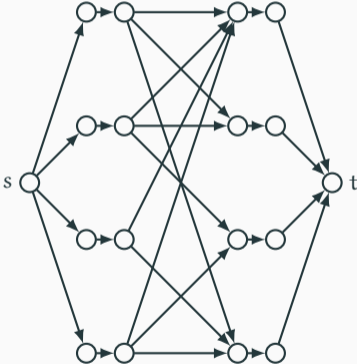


FROM VERTEX TO EDGE UNRELIABILITY IN DAGs

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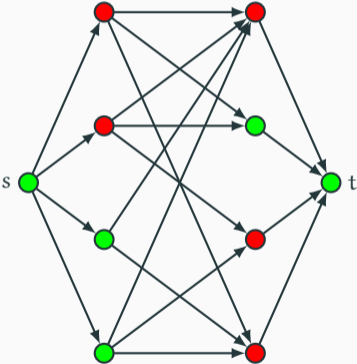


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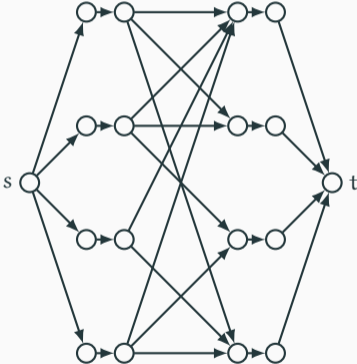


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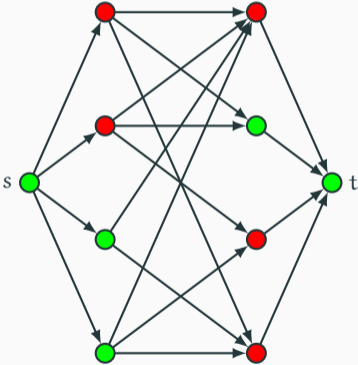


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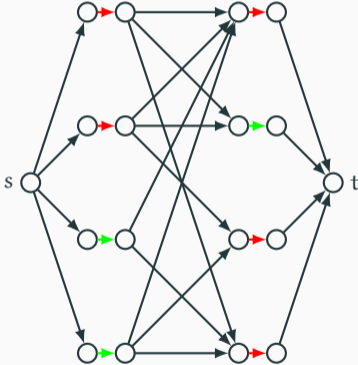


FROM VERTEX TO EDGE UNRELIABILITY IN DAGs

VERTEX UNRELIABILITY



EDGE UNRELIABILITY



Black edges have failure probability 0. Colored edges have failure probability 1/2.

CONCLUDING REMARKS

OPEN PROBLEMS

- Faster than $O(m^4 n^6 \max\{m^4, \varepsilon^{-4}\})$?

- Efficient deterministic algorithms?

This would require derandomise the [Karp-Luby](#) algorithm, which has been studied but a polynomial-time algorithm is still elusive.

- Approximation complexity for [TWO-TERMINAL RELIABILITY](#) in directed and undirected graphs?
- Improve the quasi-polynomial time algorithm for [#CFG](#) by [Gore, Jerrum, Kannan, Sweedyk, and Mahaney \(1997\)](#).

THANK YOU!

arXiv:2310.00938

THE OVERALL RUNNING TIME

We need union bounds over all edges at various points. Also they accumulate during the DP. Thus set the error $\delta := n^{-1} \min\{m^{-1}, \varepsilon\}$ in **Karp-Luby**, and we need $O(n\delta^{-2})$ samples each time. (In fact here we use a two-stage algorithm to be more efficient.) To compute the estimator, we need $O(m)$ time, so one **Karp-Luby** step takes $O(mn\delta^{-2})$ time.

One run of **Karp-Luby** succeeds with constant probability. To show that reusing samples is fine, we need a union bound over all possible scenarios in the sampling algorithm, which are exponentially many. Thus, we repeat **Karp-Luby** $O(m)$ times to achieve $\exp(-O(m))$ failure probability.

How many samples do we need for each S_i ?

On average, we need $\ell := \frac{O(mn\delta^{-2})}{n} = O(m\delta^{-2})$ samples. This in fact suffices.

At each step of the DP, we do approximate counting once and sampling ℓ times, and each sampling step, in turn, calls approximate counting $O(m)$ times.

Thus, the overall running time is

$$n \times m\ell \times m \times O(mn\delta^{-2}) = O(m^4 n^2 \delta^{-4}) = O(m^4 n^6 \max\{m^4, \varepsilon^{-4}\}).$$

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NP-HARDNESS FOR $S - t$ RELIABILITY IN DAGs

We reduce from **MINIMUM DIRECTED STEINER TREE** in DAGs with unit weights, which is **NP**-hard by reducing from **SET COVERS**.

Note that minimum Steiner trees are the extremal configurations for $S - t$ RELIABILITY. Thus, if the failure probability of edges is $1 - \exp(-O(n))$, then, drawing from the corresponding distribution, we would only see the minimizers.

Exponentially small success probability can be simulated by gadgets, such as replacing a normal edge by a path (sometimes called a stretching).